

Electronic Journal of Applied Statistical Analysis EJASA (2012), Electron. J. App. Stat. Anal., Vol. 5, Issue 1, 15 – 29 e-ISSN 2070-5948, DOI 10.1285/i20705948v5n1p15 © 2012 Università del Salento – http://siba-ese.unile.it/index.php/ejasa/index

# MODELING AND FORECASTING EXCHANGE RATE DYNAMICS IN PAKISTAN USING ARCH FAMILY OF MODELS

Shahla Ramzan<sup>\*(1)</sup>, Shumila Ramzan<sup>(2)</sup>, Faisal Maqbool Zahid<sup>(1)</sup>

<sup>(1)</sup> Department of Statistics, Government College University Faisalabad, Pakistan <sup>(2)</sup>University of Agriculture, Faisalabad, Pakistan

> Received 14 October 2010; Accepted 05 October 2011 Available online 26 April 2012

Abstract: The main objective of this paper is to provide an exclusive understanding about the theoretical and empirical working of the GARCH class of models as well as to exploit the potential gains in modeling conditional variance, once it is confirmed that conditional mean model errors present time varying volatility. Another objective is to search the best time series model among autoregressive moving average (ARMA), autoregressive conditional heteroscedasticity (ARCH), generalized autoregressive conditional heteroscedasticity (GARCH), and exponential generalized autoregressive conditional heteroscedasticity (EGARCH) to give best prediction of exchange rates. The data used in present study consists of monthly exchange rates of Pakistan for the period ranging from July 1981 to May 2010 obtained from State Bank of Pakistan. GARCH (1,2) is found to be best to remove the persistence in volatility while EGARCH(1,2) successfully overcome the leverage effect in the exchange rate returns under study.

*Keywords*: Conditional variance, Exchange rates, GARCH, EGARCH, Volatility modeling.

# 1. Introduction

In the era of globalization and financial liberalization, exchange rates play an important role in international trade and finance for a small open economy like Pakistan. This is because movements in exchange rates affect the profitability of multinationals and increase exchange exposure to enterprises and financial institutions. A stable exchange rate may help venture and financial institutions in evaluating the performance of investments, financing and hedging, and

<sup>\*</sup> E-mail: <u>shahla\_ramzan@yahoo.com</u>

thus reducing their operational risks. Fluctuations in the exchange rate may have a significant impact on the macroeconomic fundamentals such as interest rates, prices, wages, unemployment, and the level of output. The behavior of exchange rates is of crucial interest for the policy makers.

Recent developments in financial econometrics require the use of quantitative models that are able to explain the attitude of investors not only towards the expected returns and risks, but towards volatility as well. Hence, market participants should be aware of the need to manage risks associated with volatility, which requires models that are capable of dealing with the volatility of the market. Due to unexpected events like, the non-constant variance in the financial markets, and uncertainties in prices and returns, financial analysts started to model and explain the behavior of exchange rate returns and volatility using time series econometric models.

One of the most prominent tools for capturing such changing variance is the Autoregressive Conditional Heteroscedasticity (ARCH) and Generalized Autoregressive Conditional Heteroscedasticity (GARCH) models. These models were developed by Engle [1] and extended by Baillie [2] and Nelson [3]. Two important characteristics within financial time series, the fat tails and volatility clustering (or volatility pooling), can be captured by the GARCH family models. Many studies have been done to model the exchange rates of a country. Corte et al [4] used Bayesian methods for estimation and ranking of a set of empirical exchange rate models, and constructed combined forecasts based on Bayesian model averaging. Sing [5] investigated the GARCH model for a comprehensive set of weighted (export and trade) and unweighted (official and black) real exchange rate series in India. He found evidence of dimensionally weak and statistically insignificant autoregressive conditional heteroscedasticity effects as compared to GARCH effects in almost all the exchange rate series. Other studies include Baillie and Bollerslev [2], Poon and Granger [6], Arize et al [7], Fidrmuc and Horváth [8] and Rapach and Strauss [9].

The economy of Pakistan is the 27th largest economy in the world in terms of purchasing and the 45th largest in absolute dollar terms. Pakistan has been categorized as an emerging market. There are three stock exchange markets in Pakistan. The Karachi Stock Exchange (KSE) being the most liquid and the biggest in terms of market capitalization and trading volume and been awarded the title of "best performing emerging stock market of the world" in 2002 by Business Week. Jalil and Feridun [10] explained the exchange rate movements in the Pakistani foreign exchange market using the market micro structure approach, which had not been applied to date due to the unavailability of high-frequency data on the order flow for Pakistan. Mohammad [11] empirically assess the effect of Euro-Dollar Exchange rate on chosen macroeconomic variables like real output, price level, and money supply for Pakistan. Mohammad [11] applied vector autoregessive (VAR) approach to find the relation between the above and the results are evident of no significant impact of Euro and US dollar depreciation on the macroeconomic variables of Pakistan. This paper offers insights on exchange rates in Pakistan and measures the sources of volatility by using Autoregressive Conditional Heteroscedasticity (ARCH), Generalized Autoregressive Conditional Heteroscedasticity (GARCH) and Exponential General Autoregressive Conditional Heteroscedasticity (EGARCH) techniques.

The rest of this paper is categorized as follows: Section 2 presents the description of the data used in this study and some econometric methodology used for analysis. Section 3 presents statistical features of the data and results of the analysis with some discussion. Finally last section concludes the text with some concluding remarks.

## 2. Material and Methods

### 2.1 Data and Econometric methodology

For conventional analysis and volatility modeling, the data used in the present study consists of monthly average Foreign Exchange rates of Pakistan (Pak rupees per US \$). The data is obtained from State Bank of Pakistan (SBP) Karachi with sample period ranging from July 1981 to May 2010. Since the official exchange markets remain close on Sunday, there are total of 347 monthly observations. To model this data, different econometric time series models like ARMA, ARCH, GARCH, IGARCH and EGARCH are used. A short description of each is given in the following text.

## 2.2 ARMA (p,q)

In an ARMA (p, q) process, there are p autoregressive and q moving average terms. The algebraic representation of the statistical model is:

$$y_{t} = \mu + \Phi_{1}y_{t-1} + \Phi_{2}y_{t-2} + \dots + \Phi_{p}y_{t-p} + \varepsilon_{t} + \theta_{1}\varepsilon_{t-1} + \theta_{2}\varepsilon_{t-2} + \dots + \theta_{q}\varepsilon_{t-q}$$
(1)

where the intercept parameter  $\mu$  is related to the mean of  $y_t$  and errors  $\varepsilon_t$  are assumed to be uncorrelated random variables, with  $E[\varepsilon_t] = 0$  and  $var[\varepsilon_t] = \sigma_t^2$ . If this process is stationary, then it must have a constant mean  $\mu$  for all time periods.

### 2.3 Autoregressive Conditional Heteroscedasticity (ARCH) Model

The autoregressive conditional heteroscedasticity (ARCH) model is the first model of conditional heteroscedasticity. In econometrics, a model featuring autoregressive conditional heteroscedasticity considers the variance of the current error term or innovation to be a function of the actual sizes of the previous time periods' error terms. Often the variance is related to the squares of the previous innovations. Such models often called ARCH models were introduced by Engle [1] and is known as a first order autoregressive conditional heteroscedasticity ARCH(1) process.

Specifically, let  $\varepsilon_t$  denote the error terms of return residuals with respect to mean process and *iid* 

assume that  $\varepsilon_t = \sigma_t z_t$  where  $z_t \sim N(0,1)$  and the series  $\sigma_t^2$  are modeled by:

$$\sigma_t^2 = \alpha_0 + \alpha_1 \varepsilon_{t-1}^2 + \dots + \alpha_q \varepsilon_{t-q}^2 = \alpha_0 + \sum_{i=1}^q \alpha_i \varepsilon_{t-i}^2$$
(2)

Here  $\alpha_0 > 0$  and  $\alpha_i \ge o$ , i > o. An ARCH(q) model can be estimated using ordinary least squares.

#### 2.4 Generalized Autoregressive Conditional Heteroscedasticity (GARCH) Model

Unfortunately, volatility is not an easy phenomenon to predict or forecast. One class of models which have proved successful in forecasting volatility in many situations is the GARCH family

of models. The Generalized Autoregressive Conditional Heteroscedasticity (GARCH) models were propounded by Engle [1] and Baillie [2]. The GARCH (p, q) model is formulated as:

$$\sigma_{t}^{2} = \alpha_{0} + \alpha_{1}\varepsilon_{t-1}^{2} + \dots + \alpha_{q}\varepsilon_{t-q}^{2} + \beta_{1}\sigma_{t-1}^{2} + \dots + \beta_{p}\sigma_{t-p}^{2} = \alpha_{0} + \sum_{i=1}^{q}\alpha_{i}\varepsilon_{t-i}^{2} + \sum_{i=1}^{p}\beta_{i}\sigma_{t-i}^{2}$$
(3)

where p is the order of the GARCH (lagged volatility) terms, and q is the order of the ARCH (lagged squared-error) terms.

## 2.5 Integrated Generalized Autoregressive Conditional Heteroscedasticity (IGARCH) Model

Integrated Generalized Autoregressive Conditional Heteroscedasticity IGARCH is a restricted version of the GARCH model, where the sum of the persistent parameters sum up to one, and therefore there is a unit root in the GARCH process. The condition for this is:

$$\sum_{i=1}^{p} \beta_{i} + \sum_{i=1}^{q} \alpha_{i} = 1$$
(4)

#### 2.6 Exponential General Autoregressive Conditional Heteroscedastic (EGARCH) Model

The Exponential General Autoregressive Conditional Heteroscedasticity (EGARCH) model is another popular GARCH model which is introduced by Nelson [3]. In EGARCH model there is no need for the parameters to follow nonnegative restriction and another important feature of EGARCH model is that it successfully capture the leverage effect while GARCH model is not to do so. EGARCH ensures that the variance is always positive even if the parameters are negative. The EGARCH (1, 1) could be given as:

$$\ln h_t^2 = \alpha_0 + \alpha_1 \left| \frac{\varepsilon_{t-1}}{h_{t-1}} \right| + \phi \frac{\varepsilon_{t-1}}{h_{t-1}} + \beta_1 \ln h_{t-1}^2$$
(5)

where  $\phi$  captures the leverage effect (i.e. asymmetric effect).

#### 2.7 Forecasting Performance

Commonly used measures for comparison of the forecasting performance are Mean Absolute Error (MAE), Root Mean Square Error (RMSE), Mean Absolute Percentage Error (MAPE), and Theil-U inequality ([12]). The selected GARCH (1, 2) model in our study is evaluated by using all these measures for evaluating the forecasting performance. The mentioned forecast error statistics are defined as:

Mean Absolute Error = 
$$\frac{\sum_{t=T+1}^{T+k} |\hat{y}_t - y_t|}{h}$$
 (6)

Root Mean Squared Error = 
$$\sqrt{\frac{\sum_{t=T+1}^{T+k} (\hat{y}_t - y_t)^2}{h}}$$
 (7)

These two forecast error statistics depend on the scale of the dependent variable. These should be used as relative measures to compare forecasts for the same series across different models. The smaller the error, the better is the forecasting ability of that model according to a given criterion. The remaining two statistics are scale invariant. The Theil inequality coefficient always lies between zero and one, where zero indicates a perfect fit.

Mean Absolute Percentage Error = 
$$100 \times \left[\frac{\sum_{t=T+1}^{T+k} \left|\frac{\hat{y}_t - y_t}{y_t}\right|}{h}\right]$$
  
(8)
  
Theil Inequality Coefficient =  $\frac{\sqrt{\frac{\sum_{t=T+1}^{T+k} (\hat{y}_t - y_t)^2}{h}}}{\sqrt{\frac{\sum_{t=T+1}^{T+k} \hat{y}_t^2}{h}} + \sqrt{\frac{\sum_{t=T+1}^{T+k} y_t^2}{h}}$ 

#### 3. **Results and Discussion**

#### 3.1 **Descriptive statistics**

To assess the distributional properties of the exchange rate return data, various descriptive statistics are reported in Table 1.

Table 1. Summary statistics.							
Mean	Std. Dev.	Skewness	Kurtosis	Jarque-Bera	Prob.*		
0.27	0.57	2.12	8.48	691.07	< 0.001		

#### As expected for a time series of returns, the mean is close to zero. The standard deviation is also high which indicates high level of fluctuations of the exchange rate returns. There is also evidence of positive skewness, with long right tail indicating that exchange rate has non symmetric returns.

The histogram of the series in Figure 1 reports that the exchange rate returns are leptokurtic or fat tailed because of its large kurtosis value. The return series of exchange rate is non normal according to the Jarque-Bera test with *p*-value less than 0.001 (see Table 1). So the hypothesis of normality of the exchange rate returns is rejected.



#### 3.2 Tests for checking Stationarity

It is important to confirm whether the data is stationary before using the time series models like Autoregressive Integrated Moving Average (ARIMA), Autoregressive Conditional Heteroscedasticity (ARCH), Generalized Autoregressive Conditional Heteroscedasticity (GARCH) and Exponential General Autoregressive Conditional Heteroskedastic (EGARCH), which are being used in our study. Three important methods of checking stationarity of time series are graphical analysis, Correlogram and unit root test.



Figure 2. Time series plot of exchange rates series.

In Figure 2 the graphical analysis of exchange rate series shows an upward trend, suggesting that the mean of the exchange rates has been changing. This perhaps suggests that the exchange rate series is not stationary.

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
	1	1	0.990	0.990	343.12	0.000
1	11	2	0.980	-0.023	680.15	0.000
1	ı dı 🛛	3	0.969	-0.026	1010.8	0.000
I	ı d ı	4	0.958	-0.026	1334.9	0.000
	1 1	5	0.947	-0.007	1652.5	0.000
	1 1	6	0.936	-0.007	1963.4	0.000
	1 1	7	0.925	-0.001	2268.0	0.000
	1 1	8	0.914	-0.001	2566.2	0.000
	1 1	9	0.903	-0.003	2858.2	0.000
	1 1	10	0.892	-0.008	3143.9	0.000
	11	11	0.881	-0.002	3423.5	0.000
1	1 1	12	0.870	0.003	3697.2	0.000
1	1 1	13	0.859	-0.005	3965.0	0.000
	1 1	14	0.849	-0.007	4226.9	0.000
1	1 1	15	0.838	-0.007	4483.0	0.000
	1 1	16	0.827	0.002	4733.5	0.000
1	1 1	17	0.817	-0.006	4978.3	0.000
	1 1	18	0.806	-0.005	5217.6	0.000
1	11	19	0.795	-0.024	5451.1	0.000
	10	20	0.784	-0.019	5678.9	0.000
	i 🏻 i	21	0.774	0.038	5901.4	0.000
	i þi	22	0.764	0.033	6119.1	0.000
	ւի	23	0.756	0.045	6332.6	0.000
	i þi	24	0.748	0.038	6542.5	0.000
	1 <b>I</b> I	25	0.740	-0.019	6748.7	0.000

Figure 3. Correlogram of exchange rate series.

The Correlogram of the series is shown in Figure 3 which indicates a patter up to 30 lags. The autocorrelation coefficient starts with a very high value of 0.990 at lag 1 and declines very slowly towards zero. It seems that the exchange rate series is nonstationary. Table 2 presents the results of Augmented Dickey–Fuller (ADF) test and Phillips-Perron (PP) test on exchange rate data. Since the statistic value for both ADF and PP tests is greater than their corresponding critical values, so we do not reject the null hypothesis of the presence of unit root in the series and conclude that the exchange rate series is nonstationary.

To transform the nonstationary exchange rate series we calculate the exchange rate returns as

$$R_t = \ln(y_t / y_{t-1}) * 100 \tag{10}$$

The time series plot of the transformed data that is exchange rate returns series is shown in Figure 4. This plot shows that the mean of the series is now about constant. So we can assume that the series is stationary.

Augmented Dickey-	Fuller test statistic	t-Statistic	Prob.*
		1.16	0.99
Test critical values:	1% level	1.16         % level       -3.45         % level       -2.87         0% level       -2.57         Adj. t-Stat         1.20	
	5% level	-2.87	
	10% level	-2.57	
Phillips-Perron test		Adj. t-Stat	Prob.*
		1.20	0.99
Test critical values:	1% level	-3.45	
	50/1 and	-2 87	
	370 IEVEI	2.07	

Table 2. ADF and Phillip-Perron test on exchange rate series.

Furthermore, the Correlogram in Figure 5 shows no trend in exchange rate returns, hence suggesting that the exchange rate return series is stationary.



Figure 4. Time series plot of exchange rate returns

Image: 1       Image: 1       1       0.427       0.427       63.714       0.         Image: 1       Image: 1       1       2       0.175       -0.010       74.394       0.         Image: 1       Image: 1       Image: 1       3       0.108       0.045       78.477       0.         Image: 1       Image: 1       Image: 1       Image: 1       1       4       0.042       -0.023       79.099       0.         Image: 1       Image: 1       Image: 1       Image: 1       1       4       0.042       -0.023       79.099       0.         Image: 1       Image: 1       Image: 1       Image: 1       1       5       0.035       0.025       79.533       0.         Image: 1       Image: 1       Image: 1       1       6       -0.003       -0.033       79.535       0.         Image: 1       Image: 1       Image: 1       7       -0.057       -0.057       80.695       0.         Image: 1       Image: 1       Image: 1       1       9       -0.002       -0.020       80.700       0.         Image: 1       Image: 1       Image: 1       1       0.019       -0.013       80.829       0. <th></th>	
Image: 1 minimum of the system of the sys	rob
Image:	.000
I       I       I       4       0.042       -0.023       79.099       0.         I       I       I       I       5       0.035       0.025       79.533       0.         I       I       II       II       6       -0.003       -0.033       79.535       0.         II       II       II       II       II       7       -0.057       -0.057       80.695       0.         II       II       II       II       II       9       -0.002       -0.020       80.700       0.         II       II       III       III       10       -0.019       -0.013       80.829       0.         III       III       III       III       10       -0.010       82.597       0.         IIII       III       III       III       12       0.028       -0.041       82.884       0.         IIII       III       III       III       13       -0.098       84.629       0.         IIII       III       IIII       IIII       IIII       14       -0.055       0.001       85.717       0.	.000
III       III       5       0.035       0.025       79.533       0.         III       III       III       6       -0.003       -0.033       79.535       0.         III       III       III       7       -0.057       -0.057       80.695       0.         III       III       III       10       8       0.003       0.060       80.699       0.         III       III       III       9       -0.002       -0.020       80.700       0.         III       III       III       10       -0.019       -0.013       80.829       0.         IIII       III       III       11       10.070       0.100       82.597       0.         IIII       III       IIII       III       12       0.028       -0.041       82.884       0.         IIII       IIII       IIII       IIII       13       -0.069       -0.098       84.629       0.         IIII       IIII       IIII       IIIII       IIIIIIIIIIIIIIIIIIIIIIIIIIIIIIIIIIII	.000
I       I       I       I       6       -0.003       -0.033       79.535       0.         I       I       I       I       7       -0.057       -0.057       80.695       0.         I       I       I       I       I       8       0.003       0.060       80.699       0.         I       I       II       II       9       -0.002       -0.020       80.700       0.         II       II       II       II       10       -0.019       -0.013       80.829       0.         III       III       III       III       10       -0.019       -0.013       80.829       0.         IIII       III       III       III       0.070       0.100       82.597       0.         IIII       III       III       III       12       0.028       -0.041       82.884       0.         IIII       IIII       IIII       IIII       IIII       13       -0.098       84.629       0.         IIIII       IIII       IIIIIIIIIIIIIIIIIIIIIIIIIIIIIIIIIIII	.000
I       I       I       7       -0.057       -0.057       80.695       0.         I       I       I       I       8       0.003       0.060       80.699       0.         I       I       I       I       9       -0.002       -0.020       80.700       0.         I       I       I       I       10       -0.019       -0.013       80.829       0.         I       I       I       I       0.070       0.100       82.597       0.         I       I       I       I       0.070       0.100       82.597       0.         I       I       I       0.070       0.100       82.884       0.         I       I       I       13       -0.069       -0.098       84.629       0.         I       I       I       I       I       14       -0.055       0.001       85.717       0.	.000
I       I       I       I       8       0.003       0.060       80.699       0.         I       I       I       9       -0.002       -0.020       80.700       0.         I       I       I       10       -0.019       -0.013       80.829       0.         I       I       I       0.070       0.100       82.597       0.         I       I       0.070       0.100       82.597       0.         I       I       12       0.028       -0.041       82.884       0.         I       I       I       13       -0.069       -0.098       84.629       0.         I       I       I       I       14       -0.055       0.001       85.717       0.	.000
I       I       I       9       -0.002       -0.020       80.700       0.         II       II       II       10       -0.019       -0.013       80.829       0.         II       II       0.070       0.100       82.597       0.         II       II       11       0.070       0.100       82.597       0.         II       II       II       12       0.028       -0.041       82.884       0.         II       II       II       13       -0.069       -0.098       84.629       0.         II       II       II       II       III       14       -0.055       0.001       85.717       0.	.000
III       III       10       -0.019       -0.013       80.829       0.         III       III       11       0.070       0.100       82.597       0.         III       III       III       12       0.028       -0.041       82.884       0.         III       III       III       13       -0.069       -0.098       84.629       0.         III       III       III       III       III       14       -0.055       0.001       85.717       0.	.000
Image:	.000
III     III     III     III       III     III     III     IIII	.000
「「「「」」」(「」」) 「「」」(「」) 「「」」(「」) 「「」」(「」) 「」」 「」」 「」」 「」」 「」」 「」」 「」」 「」」 「」」	.000
ען ין 14 -0.055 0.001 85.717 0.	.000
	.000
	.000
ען און און אין אין און און און און און און און און און או	.000
· · · · · · · · · · · · · · · · · · ·	.000
uli uli 17 -0.070 -0.073 88.070 0.	.000
u <b>[</b> )           18 -0.066 0.013 89.684 0.	.000
	.000
· · · · · · · · · · · · · · · · · · ·	.000
uli uli 21 -0.032 -0.036 90.127 0.	.000
· · · · · · · · · · · · · · · · · · ·	.000
	.000
ים 1.968 0.002 (גער 1.968 0.	.000
ען וווין וווין 125 0.059 0.021 93.287 0.	.000

Figure 5. Correlogram of exchange rate returns.

In Table 3 the results of Augmented Dickey Fuller (ADF) test and Phillips Perron (PP) test show that the exchange rate return series is stationary. The variance is high that clearly exhibit volatility clustering, which allows us to carry on further to apply the ARCH family models.

Augmented Dickey-l	Fuller test statistic	t-Statistic	Prob.*
		-11.74	< 0.01
Test critical values:	-3.45	-3.45	
	-2.87	-2.87	
	-2.57	-2.57	
Phillips-Perron test		Adj. t-Stat	Prob.*
		-11.73	< 0.01
Test critical values:	-3.45	-3.45	
Test critical values:	-3.45 -2.87	-3.45 -2.87	

Table 3. ADF and PP test on exchange rate return series.

## 3.3 Model fitting

As the descriptive statistics given in Table 1 reflect that the distribution of the exchange rate return series does not follow a normal distribution it means volatility clustering is present. Suitable econometric modeling techniques are required for our exchange rate return series. To start with, we model the conditional mean process by autoregressive process AR(1) and moving average MA(1). Both of these processes demonstrate high correlation in residuals. Among different models applied to the data, ARMA(1, 2) appears to be relatively better fit on the basis of Akaike criterion, Schwarz criterion, and Durbin Watson statistic. The results of ARMA(1, 2) are shown in Table 4.

	Coefficient	Std. Error	t-Statistic	Prob.
С	0.27	0.05	5.66	< 0.01
AR(1)	0.44	0.05	8.36	< 0.01
MA(2)	-0.035	0.06	-0.59	0.55
R-squared	0.18	Mean depe	endent var	0.27
Adjusted R-squared	0.18	S.D. deper	ndent var	0.57
S.E. of regression	0.51	Akaike inf	o criterion	1.52
Sum squared resid	90.35	Schwarz criterion		1.55
Log likelihood	-258.41	Hannan-Quinn criter.		1.53
F-statistic	38.42	Durbin-W	atson stat	2.02
Prob(F-statistic)	< 0.01			

The correlogram of ARMA(1, 2) residuals in Figure 6 shows that none of the autocorrelation and partial autocorrelation is individually statistically significant. In other words the correlogram of both autocorrelation and partial autocorrelation give the impression that the estimated residuals are purely random. Hence, there is no need to search out for another ARMA model.

Once an appropriate ARMA model for conditional mean has been identified and estimated, the next step is to test whether the estimated errors are heteroscedastic or not. For this purpose the correlogram of squared residuals from ARMA(1, 2) is presented in Figure 7.

It shows high quantity of autocorrelation in residuals which suggest to carry on further to apply the ARCH family of models. The ARCH family of models requires the presence of 'ARCH effect' in the residuals. To test the presence of ARCH effect, we use the Lagrange Multiplier (LM) test for exchange rate returns series as suggested by Engle [1]. The results of Lagrange Multiplier test are presented in Table 5. The *p*-value indicates that there is evidence of remaining ARCH effect. So, we reject the null hypothesis of absence of ARCH effect even at 1% level of significance.

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
	10	1	-0.008	-0.008	0.0246	
1 1	1	2	0.001	0.000	0.0247	
ı <u>p</u> i	ի ին	3	0.041	0.041	0.6192	0.431
1	10	4	-0.016	-0.015	0.7038	0.703
1 1	ի դիս	5	0.032	0.031	1.0583	0.787
11	i i	6	0.013	0.012	1.1193	0.891
I <mark>C</mark> I	ן ום	7	-0.080	-0.079	3.4119	0.637
1 ]1	ի դիս	8	0.038	0.034	3.9275	0.686
1 1	1	9	0.002	0.003	3.9290	0.788
10	וםי	10	-0.058	-0.053	5.1156	0.745
ים	ן ו	11	0.095	0.090	8.3197	0.502
1 🛛 1	10	12	0.038	0.045	8.8255	0.549
<b>q</b> i	[]	13	-0.084	-0.083	11.367	0.413
1	וןי	14	-0.022	-0.038	11.544	0.483
11	1	15	-0.009	0.000	11.574	0.563
111		16	0.019	0.021	11.710	0.630
I I I	ן יםי	17	-0.055	-0.070	12.825	0.616
10	וןי	18	-0.060	-0.039	14.129	0.589
1 ]1	ון ו	19	0.033	0.038	14.531	0.629
ւիւ	i) i	20	0.028	0.016	14.825	0.674
10	10	21	-0.040	-0.031	15.410	0.696
10	101	22	-0.027	-0.030	15.672	0.737
111		23	0.014	0.007	15.749	0.784
i <mark>p</mark> i	ի դին	24	0.044	0.047	16.458	0.793
i ]i	ի հեր	25	0.038	0.046	16.997	0.809

Figure 6. Correlogram of the residuals of ARMA(1, 2) model.

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
ı 🗖		1	0.414	0.414	59.537	
· 🗖	10	2	0.216	0.054	75.837	
ı <u>p</u> i	<b>(</b> )	3	0.027	-0.096	76.097	0.000
1	ון ו	4	0.015	0.029	76.175	0.000
10	10 1	5	-0.033	-0.037	76.562	0.000
10	10 1	6	-0.044	-0.029	77.259	0.000
IC I	10	7	-0.078	-0.049	79.420	0.000
Q ·	10	8	-0.089	-0.044	82.250	0.000
1	ի հեր	9	-0.024	0.050	82.448	0.000
i pi	ן ו	10	0.063	0.082	83.858	0.000
i <mark>p</mark> i	I]I	11	0.082	0.021	86.242	0.000
i pi	10	12	0.043	-0.025	86.913	0.000
1 1	10 1	13	-0.002	-0.029	86.914	0.000
10	וםי	14	-0.053	-0.055	87.941	0.000
10	1	15	-0.057	-0.018	89.107	0.000
10	I)I	16	-0.030	0.017	89.439	0.000
1	I)I	17	-0.015	0.008	89.524	0.000
1 1	I)I	18	-0.000	0.021	89.524	0.000
ւիւ	ի հեր	19	0.037	0.046	90.025	0.000
1 1	וןי	20	0.005	-0.046	90.035	0.000
1	10 -	21	-0.022	-0.049	90.210	0.000
1		22	-0.015	0.002	90.297	0.000
ı <mark>p</mark> ı	ן ו	23	0.067	0.096	91.981	0.000
10	[]	24	-0.034	-0.101	92.420	0.000
11	l ili	25	-0.021	0.017	92.579	0.000

Figure 7. Correlogram of square residuals of ARMA (1, 2) model.

F-statistic	35.69	Prob. F(2,340)	< 0.01
Obs*R-squared	59.51	Prob. Chi-Square(2)	< 0.01

Table 5. ARCH LM test on ARMA(1, 2) residuals.

Therefore the conditional heteroscedasticity in estimated ARMA errors is modeled using GARCH(1, 2) specification. The results are shown in Table 6. The estimated coefficient of both conditional mean equation and conditional variance equation of GARCH(1, 2) model are highly significant. Therefore GARCH(1, 2) model can be cited as the suitable model on the basis of AIC and BIC criteria given in Table 6 and correlogram of the squared residuals from GARCH(1, 2) in Figure 8.

Table 6. Estimation Result of GARCH(1, 2) model.								
	Coefficient	Std. Error	z-Statistic	Prob.				
С	0.0009	7.58E-05	12.0669	< 0.001				
AR(1)	-0.9511	0.0048	-197.6292	< 0.001				
MA(2)	-0.9028	0.0107	-84.0787	< 0.001				
	Variance Equation							
С	0.0668	0.0155	4.3060	< 0.001				
RESID(-1)^2	2.1829	0.7702 2.8343		0.005				
GARCH(-1)	0.1446	0.0270	5.3479	< 0.001				
GARCH(-2)	-0.0237	0.0072	-3.2997	0.001				
GED PARAMETER	0.504927	0.040090	0.040090 12.59483					
R-squared	0.5936	Mean depen	dent var	0.0002				
Adjusted R-squared	0.5851	S.D. depend	ent var	0.9750				
S.E. of regression	0.6280	Akaike info	criterion	0.6541				
Sum squared resid	131.7403	Schwarz crit	terion	0.7438				
Log likelihood	-103.8500	Hannan-Qui	nn criter.	0.6898				
F-statistic	69.6902	Durbin-Wat	son stat	2.5448				
Prob(F-statistic)	< 0.001							

Once the heteroscedasticity is modeled using appropriate GARCH model, next we check whether the GARCH(1, 2) has adequately captured the persistence in volatility and there is no ARCH effect left in the residuals from the selected models. The ARCH LM test is conducted for this purpose. The results of LM test given in Table 7 indicate that the residuals do not show any ARCH effect. Hence, GARCH(1, 2) is found to be reasonable to remove the persistence in volatility.

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
		1	-0.048	-0.048	0.7942	
		2		-0.128	6.2526	
10 1	101	3	-0.060	-0.075	7.5169	0.006
d I		4	-0.084	-0.111	9.9868	0.007
1 🛛 1		5	0.036	0.005	10.427	0.015
1 1	1 101	6	-0.004	-0.034	10.433	0.034
ı <b>d</b> ı	n <u>e</u> li	7	-0.066	-0.079	11.976	0.035
ւի	i]i	8	0.041	0.020	12.579	0.050
ı <b>j</b> ı	i)i	9	0.033	0.019	12.969	0.073
ı <b>d</b> ı	n <u>d</u> i	10	-0.072	-0.077	14.809	0.063
· 🗖 ·		11	0.126	0.121	20.469	0.015
u (	(1	12	-0.045	-0.039	21.176	0.020
1 1	i)i	13	-0.008	0.013	21.198	0.031
1	1 1	14	0.011	0.002	21.243	0.047
ւի	լ թ	15	0.047	0.079	22.038	0.055
ı <b>p</b> i	ı <u> </u> ı	16	0.039	0.038	22.574	0.068
1 1	i <b>)</b> i	17	-0.000	0.021	22.574	0.094
el i	ի պիս	18	-0.095	-0.058	25.840	0.056
ı þi	ip	19	0.084	0.095	28.413	0.040
111		20	0.022	0.006	28.587	0.054
1	ı <u> </u> ı	21	-0.007	0.034	28.607	0.072
1		22	-0.002	-0.013	28.609	0.096
u <b>t</b> i		23	-0.033	0.004	29.017	0.114
ւի	ի հեր	24	0.064	0.053	30.516	0.106
ı <b>p</b> i	ի դի	25	0.030	0.035	30.840	0.127
I	I I	I				

Figure 8. Correlogram of squared residuals from GARCH (1, 2) model.

Table 7. ARCH LM test on GARCH (1, 2) residuals.

F-statistic	0.4801	Prob. F(1,339)	0.4889
Obs*R-squared	0.4822	Prob. Chi-Square(1)	0.4874

Normality test of standardized squared residuals, however suggests that residuals are positively skewed due to the leverage effect. EGARCH is used to test the leverage effect that successfully captures the asymmetry. The results on asymmetric conditional variance for exchange rate return series are reported in Table 8. The GARCH(1, 2) parameter is significant, showing high degree of persistence. Coefficient of the asymmetric function is positive and significantly different from zero means there is no leverage effect left.

	Coefficient	Std. Error	z-Statistic	Prob.	
С	0.0007	0.0001	7.0329	< 0.001	
AR(1)	-0.9591	0.0049	-196.6726	< 0.001	
MA(2)	-0.9199	0.0088	-104.0558	< 0.001	
Variance Equation					
C(4)	0.0374	0.0266	1.4036	0.1604	
C(5)	-0.0881	0.0480	-1.8373	0.0662	
C(6)	0.3935	0.0805	4.8870	< 0.001	
C(7)	0.7996	0.2158	3.7054	< 0.001	
C(8)	0.2148	0.2141	1.0030	0.3159	
GED PARAMETER	0.4692	0.0411	11.4232	< 0.01	

Table 8. Estimation Result of EGARCH (1, 2).

#### 3.4 Forecast Analysis

Forecast performance of the fitted GARCH(1, 2) model of exchange rate returns is investigated through mean absolute error (MAE), root mean square error (RMSE), mean absolute percent error (MAPE), and Theil inequality coefficient. The results are shown in Table 9..

Root Mean Square Error	0.6921	
Mean Absolute Error	0.3257	
Mean Abs. Percent Error	390.29	
Theil Inequality Coefficient	0.4801	
Bias Proportion	0.0011	
Variance Proportion	0.0019	
Covariance Proportion	0.9971	

Table 9. Forecasting performance of exchange rate returns.

The value of Theil inequality is 0.4801 indicating the model is a better fit. The bias proportion and variance proportion are close to zero. The value of covariance proportion is nearly one. We can say that this model is good for forecasting purpose along with capturing the volatility and the leverage effects.

## 4. Conclusion

This paper focuses on building a model for the exchange rate of Pakistan using time series methodology. As the financial time series like exchange rate may possess volatility, an attempt is made to model this volatility using ARCH and GARCH models. Monthly average foreign exchange rates of Pakistan for the period ranging from July 1981 to May 2010 are used for this purpose. First of all, the stationarity of the exchange rate series is examined using graphical analysis, correlogram and unit root test which showed the series as nonstationary. To make the exchange rate series stationary, the exchange rates are transformed to exchange rate returns. Then ARMA(1, 2) model is fitted to the data. To capture the volatility, GARCH (1, 2) model is used. This model is further having leverage effect. This leverage effect is captured using an EGARCH model. Finally, the forecast performance is measured using different measures like MAE, RMSE and MAPE etc. The GARCH family of models captures the volatility and leverage effect in the exchange rate returns and provides a model with fairly good forecasting performance.

## References

- [1]. Engle, R.F. (1982). Autoregressive Conditional Heteroscedasticity with estimates of variance of United Kingdom inflation. *Econometrica*, 50(4), 987-1007.
- [2]. Baillie, R.T., Bollerslev, T. (2002). The massage in daily exchange rates: A conditional variance tale. *Journal of Business and Economic Statistics*, 20(1), 60-68.
- [3]. Nelson, D.B. (1991). Conditional heteroscedasticity in asset returns: A new approach. *Econometrica*, 59(2), 347-370.
- [4]. Corte, P.D., Sarno, L., Tsiakas, I. (2008). An Economic Evaluation of Empirical Exchange Rate Models. *The review of financial studies*, 22(9), 3491-3530.
- [5]. Singh, T. (2002). On the GARCH Estimates of Exchange Rate Volatility in India. *Applied Economics Letters*, 9, 391-395.
- [6]. Poon, S.H., Granger, C.W.J. (2003). Forecasting financial market volatility: A review. *Journal of Economic Literature*, 41(2), 478-539(62).
- [7]. Arize, A.C., Osang, T., Slottje, D. J. (2006). Exchange-rate volatility in Latin America and its impact on foreign trade. *International Review of Economics & Finance*, 17(1), 33-44.
- [8]. Fidrmuc, J., Horváth, R. (2007). Volatility of exchange rates in selected new EU members: Evidence from daily data. *Economic Systems*, 32(1), 103-118.
- [9]. Rapach, D.E., Strauss, J.K. (2008). Structural breaks and GARCH models of exchange rate volatility. *Journal of Applied Econometrics*, 23(1), 65-90.
- [10]. Jalil, A., Feridun, M. (2010). Explaining exchange rate movements: an application of the market microstructure approach on the Pakistani foreign exchange market. *The Journal of Developing Areas*, 44(1), 255-265.
- [11]. Mohammad, S.D. (2010). The Euro Dollar Exchange Rates & Pakistan Macroeconomics Dynamics. *European Journal of Scientific Research*, 42(1), 6-15.
- [12]. Abrosimova, N., Dissanaike, G., Linowski, D. (2005). Testing the weak-form Efficiency of the Russian Stock Market. http://papers.ssrn.com/sol3/papers.cfm?abstract\_id=302287

This paper is an open access article distributed under the terms and conditions of the <u>Creative Commons</u> Attribuzione - Non commerciale - Non opere derivate 3.0 Italia License.